# CIS 625 PROBLEM SET 1 

Professor Kearns

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## Problem 1.

Algorithm. Define a learning algorithm $L$ such that $L$ takes in inputs of the form $\left\langle\left(x_{1}, x_{2}, \ldots, x_{n}\right), c\left(x_{1}, x_{2}, \ldots, x_{n}\right)\right\rangle$, which we denote as $\langle x, c(x)\rangle$ and parameters $\epsilon, \delta$. We wish to output the tightest fit rectangle of $n$ dimensions.
To do this, we examine all $m$ inputs and select the smallest and largest value for each of $n$ dimensions that is positively labeled. We return the rectangle defined by the $2 n$ values that determine the min and max value (edges) of the $n$ dimensions of the rectangle.
Runtime. We see $m$ samples. For each sample, we iterate over each dimension to calculate the tightest fit on that dimension, giving us a $O(m n)$ runtime.
Error Analysis. The error analysis proceeds in a similar manner as the error analysis of PAC learning axis aligned rectangles of 2 dimensions. Consider if the points are selected to some distribution $\mathcal{D}$.
First observe that the hypothesis we return is indeed an axis aligned rectangle and that the rectangle of tightest fit is always contained within the target rectangle $R$. We define our empirical rectangle as $R^{\prime}$ and the error between the two rectangles as $R \Delta R^{\prime}=R-R^{\prime}$. Notice that we can split the error region $R \Delta R^{\prime}$ into $2 n$ strips. In the 2 dimensional case, it is the top, left, right, and bottom strips. We do this for the rectangle of $n$ dimension. We want to guarantee that the weight of all strips under $\mathcal{D}$ is at most $\epsilon$. Consider the strip with weight at most $\frac{\epsilon}{2 n}$. If each strip has weight greater than $\frac{\epsilon}{2 n}$, then our total error will exceed $\epsilon$. But a strip will only have weight greater than $\frac{\epsilon}{2 n}$ if there is no point in that region.
The probability that there is no such point landing in the region of a strip with weight $\frac{\epsilon}{2 n}$ is $\left(1-\frac{\epsilon}{2 n}\right)$. Since every point is selected independently via $m$ draws, the probability that a single strip does not have any point landing in that region is $\left(1-\frac{\epsilon}{2 n}\right)^{m}$. Using union bound, we can upper bound the probability that no point lands in any of the strips as:

$$
\sum_{i=1}^{2 n}\left(1-\frac{\epsilon}{2 n}\right)^{m}=2 n\left(1-\frac{\epsilon}{2 n}\right)^{m}
$$

So the probability that our error is greater than epsilon is at least $2 n\left(1-\frac{\epsilon}{2 n}\right)^{m}$. We want this to be less than $\delta$. Now we solve for $m$ :

$$
\begin{aligned}
2 n\left(1-\frac{\epsilon}{2 n}\right)^{m} & \leq \delta \\
-\frac{\epsilon m}{2 n} & \leq \ln \left(\frac{\delta}{2 n}\right) \\
m & \geq \frac{2 n}{\epsilon} \ln \left(\frac{2 n}{\delta}\right)
\end{aligned}
$$

Since $\mathcal{D}$ was arbitrary, this analysis must work for any distribution.

## Problem 2.

## Hypothesis/learning algorithm:

Our algorithm is simply to fit the smallest interval around each contiguous set of positive points. Let $h$ be the hypothesis we output at the end of our algorithm. We construct $h$ in $O(m \log m)$ time as follows:
(1) Sort the points given by the sample in increasing order. Since we are given a finite number of points, there must be a smallest and largest point value.
(2) Iterating from smallest point to right, keep track of the first positive example and then the last positive example that occurs before a negative example.
(3) Repeat on the next positive example.
(4) Add this interval to $h$ so that at the end, $h=\bigcup_{i=1}^{n^{*}}\left[a_{i}, b_{i}\right]$

## Analysis:

There are two places where our hypothesis can make mistakes: It can mislabel the ends of intervals as negative, and it can also mislabel some negative regions as positive; namely the negative intervals (between positive intervals) from which we don't see any examples. Note that there are $3 d-1$ such bad regions ( $2 d$ ends of positive intervals, $d-1$ sandwiched negative intervals).

We can perform a similar geometric analysis as in the case of rectangles. Consider regions of weight $\frac{\epsilon}{3 d}$ from the left and right ends of each of the positive intervals; there can be (at most) $2 d$ of these. Also, consider any sandwiched negative intervals with weight $\geq \frac{\epsilon}{3 d}$; there can be at most $d-1$ of these. As long as our sample hits each of these $3 d-1$ target intervals at least once, each of our error regions will have probability mass $<\frac{\epsilon}{3 d}$, so the total error region will have probability mass $<\frac{(3 d-1) \epsilon}{3 d}<\epsilon$. Our hypothesis is good as long as we hit each of these target intervals at least once.

Define the following events:

$$
\begin{aligned}
B_{i}(i \in[3 d-1]) & :=\text { the sample misses the } i \text { th target interval } \\
B_{i}^{j}(j \in[m]) & :=\text { the } j \text { th example misses the } i \text { th target interval } \\
B & :=\text { the sample misses at least one of the target intervals } \\
G & :=\text { the good event, where we hit all target intervals }
\end{aligned}
$$

We can then compose these events as follows:

$$
\begin{aligned}
\forall i, j: B_{i}^{j} & \leq 1-\frac{\epsilon}{3 d} \\
\forall i: B_{i} & =\prod_{j \in[m]} B_{i}^{j} \leq\left(1-\frac{\epsilon}{3 d}\right)^{m} \\
\Longrightarrow B & \leq \sum_{i \in[3 d-1]} B_{i} \leq 3 d\left(1-\frac{\epsilon}{3 d}\right)^{m} \\
\Longrightarrow G & =1-B \geq 1-3 d\left(1-\frac{\epsilon}{3 d}\right)^{m}
\end{aligned}
$$

Finally, we set $1-\delta>G$ and find that it suffices for

$$
m>\frac{3 d}{\epsilon} \log \frac{3 d}{\delta}
$$

## Problem 3.

There are many answers, but here is one such example. Notice that the rectangles are distinct and maximal, so there is not a unique target rectangle that we can analyze.


## Problem 4.

a) We wish to show that if monotone $m$-term DNF is PAC learnable by some learning algorithm that runs polynomial in $\left(\frac{1}{\epsilon}, \frac{1}{\delta}\right)$, then general $m$-term DNF is PAC learnable. We proceed with a reduction from general $m$-term DNF to monotone $m$-term DNF.

Suppose that monotone $m$-term DNF is PAC learnable with some learning algorithm $L_{m}$. Define the following mapping

$$
f:\{0,1\}^{n} \rightarrow\{0,1\}^{2 n}
$$

Such that $f$ maps input $\left(x_{1}, x_{2}, \ldots, x_{n}\right) \rightarrow\left(x_{1}, x_{2} \ldots, x_{n}, x_{n+1}, \ldots, x_{2 n}\right)$ where the $x_{n+i}$ term represents the negated literal of the $x_{i}$ term $\left(\neg x_{i}\right)$ in the input.

For example, if $n=2$, then $f(1,0)=(1,0,0,1)$.
Now we define a mapping $g$ from the set of $m$-term DNFs over $n$ variables to the set of monotone $m$-term DNFs over $2 n$ variables, which replaces any $\neg x_{i}$ with $x_{n+i}$. For example, if $n=3, g\left(x_{1} \neg x_{2} \vee \neg x_{3} \vee x_{1} x_{2}\right)=x_{1} x_{5} \vee x_{6} \vee x_{1} x_{2}$. Clearly $g$ is a bijection as for any monotone DNF, we can map it back to a general DNF via $g^{-1}$ that sends any instance of a literal to itself or the negated version (ie it replaces $x_{j}$ with $\neg x_{j-n}$ for $j>n$ given the size of $n$ ). Notice that these mappings are linear in $O(m n)$.

Let $L$ be a learning algorithm for $m$-term DNF, which consists of running an instance of $L_{m}$, and for each sample point, $\{x, c(x)\}, L$ supplies $L_{m}$ with $\{f(x), c(x)\}$. After processing the entire sample, $L$ outputs $g^{-1}(h)$, where $h$ is the hypothesis outputted by $L_{m}$ and $g^{-1}$ is the inverse of $g$ as defined above. We claim that $L$ is a PAC learning algorithm for $m$-term DNF, by $m$-term DNF.

To see why, suppose that we are given some target $m$-term DNF $c:\{0,1\}^{n} \rightarrow$ $\{0,1\}$, some sampling distribution $D$ over $\{0,1\}^{n}$, and some $\delta, \epsilon$. We claim that for any $x \in\{0,1\}^{n}$ and $m$-term DNF, $h, h(x)=(g(h))(f(x))$. To see why, suppose

$$
\begin{aligned}
& x=\left(z_{1}, z_{2}, \ldots, z_{n}\right) \text { and let } f(x)=\left(y_{1}, y_{2}, \ldots, y_{n}, \ldots y_{2 n}\right) . \\
h= & x_{a_{1}} \ldots x_{a_{k}} \neg x_{a_{k+1}} \ldots \neg x_{a_{k+j}} \vee x_{b_{1}} \ldots x_{b_{l}} \neg x_{b_{l+1}} \ldots \neg x_{b_{l+m}} \vee x_{c_{1} \ldots x_{c_{p}} \neg x_{c_{p+1}} \ldots \neg x_{c_{p+q}}} \\
h(x)= & \Longleftrightarrow h\left(z_{1}, \ldots, z_{n}\right)=1 \\
& \Longleftrightarrow z_{a_{1}} \ldots z_{a_{k}} \neg z_{a_{k+1}} \ldots \neg z_{a_{k+j}} \vee z_{b_{1} \ldots} \ldots z_{b_{l}} \neg z_{b_{l+1} \ldots \neg z_{b_{l+m}}} \vee z_{c_{1} \ldots z_{c_{p}} \neg z_{c_{p+1} \ldots} \ldots \neg z_{c_{p+q}}=1} \\
& \Longleftrightarrow y_{a_{1} \ldots y_{a_{k}}} y_{a_{k+1}+n \ldots} \ldots y_{a_{k+j}+n} \vee y_{b_{1} \ldots} \ldots y_{b_{l}} y_{b_{l+1}+n} \ldots y_{b_{l+m}+n} \vee y_{c_{1} \ldots y_{c_{p}}} y_{c_{p+1}+n} \ldots y_{c_{p+q}+n}=1 \\
& \Longleftrightarrow(g(h))\left(y_{1}, \ldots, y_{2 n}\right)=1 \\
& \Longleftrightarrow(g(h))(f(x))=1
\end{aligned}
$$

From the observation that $h(x)=(g(h))(f(x))$, we observe that running $L$ with target $c$ and distribution $D$ is equivalent to running $L_{m}$ with target $g(c)$ and some distribution $D^{\prime}$ over $\{0,1\}^{2 n}$ (which is simply the $D$ under $f$ ). Therefore by the PAC property of $L_{m}$, we may choose sample size $m$ which polynomialy varies with $\frac{1}{\delta}, \frac{1}{\epsilon}$, and see that with probability at least $1-\delta$,

$$
\begin{aligned}
\varepsilon_{D}\left(g^{-1}(h)\right) & =\operatorname{Pr}_{x \sim D}\left[c(x) \neq g^{-1}(h)(x)\right] \\
& =\operatorname{Pr}_{x \sim D}[(g(c))(f(x)] \neq(h)(f(x))) \\
& =\operatorname{Pr}_{x \sim D^{\prime}}[(g(c))(x) \neq(h)(x)]<\epsilon
\end{aligned}
$$

We now may observe that this above argument generalizes to other classes of boolean formulae and their monotone counterpart. In particular, we see that if $\mathcal{C}$ is some boolean class of (finite) functions, then the $g$ described above, is a map from $\mathcal{C}$ to its monotone-counterpart. By replicating the above analysis, we may conclude that PAC-learnability of the monotone counterpart by itself, implies PAC-learnablility of $\mathcal{C}$ by itself.
b) We wish to show that if read-once $m$-term DNF is PAC learnable by itself, then $m$-term DNF is PAC learnable by itself. Suppose that read-once $m$-term DNF is PAC learnable by itself with some learning algorithm $L_{r}$. Define a mapping

$$
f:\{0,1\}^{n} \rightarrow\{0,1\}^{m n}
$$

By sending $\left(x_{1}, \ldots, x_{n}\right) \rightarrow\left(x_{1}, \ldots, x_{n}, x_{1}, \ldots, x_{n}, x_{1}, \ldots, x_{n}\right)$. For example, if $n=3$, then $f(1,1,0)=(1,1,0,1,1,0,1,1,0)$. Now define a mapping $g$ from the set of $m$-term DNF's over to $n$ variables to the set of read-once $m$-term DNF's over $m n$ variables, which replaces any $x_{j}$ in the second conjunction with $x_{j+n}$ and any $x_{k}$ in the third conjunction with $x_{k+2 n}$. For example, if $n=3, g\left(x_{1} \neg x_{2} \vee \neg x_{3} \vee x_{1} x_{2}\right)=$ $x_{1} \neg x_{2} \vee \neg x_{6} \vee x_{7} x_{8}$. Notice that these reductions are linear in $O(m n)$.

Finally, let $L$ be a learning algorithm for $m$-term DNF, which consists of running an instance of $L_{r}$, and for each sample point, $\{x, c(x)\}, L$ supplies $L_{r}$ with $\{f(x), c(x)\}$. After processing the entire sample, $L$ outputs $g^{-1}(h)$, where $h$ is the output model of $L_{r}$ and $g^{-1}$ is the inverse of $g$ (i.e it replaces each $x_{j}$ with $\left.x_{j \bmod n}\right)$. We claim that $L$ is a PAC learning algorithm for $m$-term DNF, by $m$-term DNF.

To see why, suppose that we are given some target $m$-term DNF $c:\{0,1\}^{n} \rightarrow$ $\{0,1\}$, some sampling distribution $D$ over $\{0,1\}^{n}$, and $\delta, \epsilon$. We claim that for any $x \in\{0,1\}^{n}$ and $m$-term DNF, $h, h(x)=(g \circ h)(f(x))$. To see why, suppose $x=\left(z_{1}, z_{2}, \ldots, z_{n}\right)$ and $f(x)=\left(y_{1}, y_{2}, \ldots, y_{m n}\right)$.

$$
\begin{aligned}
h & =x_{a_{1} \ldots} \ldots x_{a_{k}} \neg x_{a_{k+1}} \ldots \neg x_{a_{k+j}} \vee x_{b_{1}} \ldots x_{b_{l}} \neg x_{b_{l+1}} \ldots \neg x_{b_{l_{l+m}}} \vee x_{c_{1} \ldots x_{c_{p}} \neg x_{c_{p+1}} \ldots \neg x_{c_{p+q}}} \\
h(x) & =1 \\
& \Longleftrightarrow h\left(z_{1}, \ldots, z_{n}\right)=1 \\
& \Longleftrightarrow z_{a_{1}} \ldots z_{a_{k}} \neg z_{a_{k+1}} \ldots \neg z_{a_{k+j}} \vee z_{b_{1} \ldots} \ldots z_{b_{l}} \neg z_{b_{l+1}} \ldots \neg z_{b_{l+m}} \vee z_{c_{1} \ldots} \ldots z_{c_{p}} \neg z_{c_{p+1}} \ldots \neg z_{c_{p+q}}=1 \\
& \Longleftrightarrow y_{a_{1} \ldots} \ldots y_{a_{k}} \neg y_{a_{k+1}} \ldots \neg y_{a_{k+j}} \vee \cdots \vee y_{c_{1}+2 n} \ldots y_{c_{p}+2 n} \neg y_{c_{p+1}+2 n} \ldots \neg y_{c_{p+q}+2 n}=1 \\
& \Longleftrightarrow(g(h))\left(y_{1}, \ldots, y_{m n}\right)=1 \\
& \Longleftrightarrow(g(h))(f(x))=1
\end{aligned}
$$

From the observation that $h(x)=(g(h))(f(x))$, we observe that running $L$ with target $c$ and distribution $D$ is equivalent to running $L_{m}$ with target $g(c)$ and some distribution $D^{\prime}$ over $\{0,1\}^{m n}$ (which is simply the $D$ under $f$ ). Therefore by the PAC property of $L_{m}$, we may choose sample size $m$ which polynomially varies with $\frac{1}{\delta}, \frac{1}{\epsilon}$, and see that with probability at least $1-\delta$,

$$
\begin{aligned}
\varepsilon_{D}\left(g^{-1}(h)\right) & =\operatorname{Pr}_{x \sim D}\left(c(x) \neq g^{-1}(h)(x)\right) \\
& =\operatorname{Pr}_{x \sim D}((g(c))(f(x)) \neq(h)(f(x))) \\
& =\operatorname{Pr}_{x \sim D^{\prime}}((g(c))(x) \neq(h)(x))<\epsilon
\end{aligned}
$$

We now make a brief remark on the above analysis. We have thus far assumed that any variable $x_{j}$ will not appear in the a single conjunction more than once. If we relax this assumption and allow $x_{j}$ and $\neg x_{j}$ to appear in a single term of a $m$-term DNF, then we can trivially modify our $f$ and $g$ to preserve the analysis. In particular we may instead make $f$ map into $\{0,1\}^{2 m n}$, by $f\left(x_{1}, \ldots, x_{n}\right)=\left(x_{1}, . ., x_{n}, x_{1}, \ldots, x_{n}\right)$. Accordingly, define $g$ to map $x_{j}$ and $\neg x_{j}$ to $x_{j+n}$ and $x_{j+2 n}$ in the second term, take $x_{j}$ and $\neg x_{j}$ to $x_{j+3 n}$ and $x_{j+4 n}$ in the third term and map $\neg x_{j}$ to $x_{j+5 n}$ in the first term.

Finally, we can generalize our entire analysis to other boolean classes. In particular, suppose $\mathcal{C}$ is some boolean class for any single variables $x_{j}$ occurs at most some $k$ times. If the read-once counterpart of $\mathcal{C}$ is PAC learnable, we may take some $g$ that maps repeated instances of $x_{j}$ in a formula in $\mathcal{C}$ to distinct $x_{j+\alpha n}$ and $x_{j+\beta n}$. Then by choosing the appropriate $f$, we can replicate the above analysis and conclude that PAC learnability of the read-once counterpart implies PAC learnability of $\mathcal{C}$.

## Problem 5.

Recall that if $L$ is a PAC learning algorithm, then for any distribution $\mathcal{D}$ and inputs $\epsilon, \delta, m, L$ returns a hypothesis $h$ such that $\varepsilon(h)<\epsilon$ with probability $1-\delta$ using $O($ poly $(1 / \epsilon, 1 / \delta, m, n)$ i.i.d. samples from $\mathcal{D}$ in $O(p o l y(1 / \epsilon, 1 / \delta, m, n)$ time.

Note that this above definition actually isn't complete in this specific case. We need to add the caveat that the examples in our samples are labeled by (i.e. consistent with) a DNF with $m$ terms. When this is not the case, we do not get any guarantees from PAC. $L$ could either crash and burn, or worse, output an arbitrarily (bad) hypothesis.

Note also that every $m$-term DNF can be written as an equivalent $m^{\prime}$ DNF for all $m^{\prime}>m$. There are several ways to do this. For example, we could pad by taking some term in the original DNF and repeat it multiple times (since $T_{1} \vee \cdots \vee T_{1} \Longleftrightarrow T_{1}$ ). As long as we
pass in some $m^{\prime}$ to $L$ which is larger than $m$ (which we do not know a priori), we have PAC guarantees.

One trivial approach is to just guess $m^{\prime}=3^{n}$ (since every DNF is equivalent to a DNF with only distinct terms, and there are $3^{n}$ possible distinct terms). Then $L$ is guaranteed to output a consistent hypothesis. But $m^{\prime}$ could be exponentially larger than $m$, so our algorithm would run in exponential time with respect to $m$ as $L$ depends polynomially on $m^{\prime}$. This fails the PAC runtime efficiency conditions.

It remains to show how we can be smarter about choosing our $m^{\prime}$ to still guarantee polynomial runtime in $m$. This is where the "doubling" trick comes in. We can search for $m^{\prime}$ by guessing it in powers of two, and running $L$ on each $m^{\prime}=2^{i}$ until we're satisfied with its hypothesis. Note that, if we are able to perfectly detect when $L$ gives us a "good" hypothesis or always return a hypothesis with error $<\epsilon$, then we simply stop at $m^{\prime}=\min _{i}\left\{2^{i}: 2^{i}>m\right\}$; so that $m^{\prime}<2 * m$, satisfying the "polynomial in $m$ " requirement.
However, because we only have a PAC algorithm, we are only able to produce a hypothesis with error $<\epsilon$ with probability $1-\delta$. So the goal now is to find a way to detect if the hypothesis produced by $L$ is good with high probability.
To do this, we use Chernoff-Hoefding bounds! The idea will be to test every hypothesis we get with a newly drawn sample and then stop at the first hypothesis which gives us a "good enough" test error.
We implement this idea in the following algorithm $L$.
$\mathbf{L}^{\prime}(\epsilon, \delta, \mathcal{D}):$
(1) Initialize $i=0$.
(2) Training: Run $L\left(\epsilon / 3, \frac{1}{3 * 2^{c}}, m=2^{i}\right)$ to get hypothesis $h_{i}$. If $2^{i}>3^{n}$, return $h_{i}$.
(3) Testing:
(a) Draw a sample $S_{i} \sim D$ of size $p=\frac{5}{\epsilon^{2}} \ln \left(\max \left\{3 * 2^{c}, \frac{2 n}{\delta}\right\}\right)$ (see analysis for why).
(b) If $\hat{\epsilon}_{S_{i}}\left(h_{i}\right)<\frac{2 \epsilon}{3}$, return $h_{i}$. Otherwise, repeat from (2) with $i=i+1$
$c$ is defined to be the largest dependence that $L$ has on $m$, across sample size and runtime. For example if the runtime of $L$ is $O\left(m^{3}\right)$ and sample complexity is $O\left(m^{4}\right)$, then $c=4$.

## Correctness:

We'll first show that $L^{\prime}$ is indeed probabilistically approximately correct. Let $h^{\prime}$ denote the hypothesis returned. And let $B$ denote the bad event:

$$
B:=\left[\epsilon\left(h^{\prime}\right)>\epsilon\right]
$$

When can this event occur? Well, it's the union of $B_{i}$, defined as the event that $h_{i}$ is bad, and $h^{\prime}=h_{i}$ (i.e. we accepted $h_{i}$ ). Therefore, we can upperbound $\operatorname{Pr}[B]$ by the union bound.

$$
\operatorname{Pr}[B] \leq \sum_{i=1}^{n \log 3} \operatorname{Pr}\left[B_{i}\right]
$$

Therefore, it suffices to upper bound every $\operatorname{Pr}\left[B_{i}\right]$ by $\operatorname{Pr}\left[B_{i}\right] \leq \frac{\delta}{n \log 3}<\frac{\delta}{2 n}$ Now, $\operatorname{Pr}\left[B_{i}\right]$ we can (loosely) upper bound by $\operatorname{Pr}\left[\right.$ accept $\left.h_{i} \mid \epsilon\left(h_{i}\right)>\epsilon\right]$. Formally, we see that

$$
\operatorname{Pr}\left[B_{i}\right] \leq \operatorname{Pr}_{S_{i}}\left[\epsilon-\hat{\epsilon}_{S_{i}}\left(h_{i}\right)>\epsilon / \underset{6}{\left.2 \mid \epsilon\left(h_{i}\right)>\epsilon\right]} \leq \operatorname{Pr}_{S_{i}}\left[\epsilon\left(h_{i}\right)-\hat{\epsilon}_{S_{i}}\left(h_{i}\right)>\epsilon / 3\right]\right.
$$

By Chernoff-Hoefding bounds, we have:

$$
\operatorname{Pr}_{S_{i}}\left[\epsilon\left(h_{i}\right)-\hat{\epsilon}_{S_{i}}\left(h_{i}\right)>\epsilon / 2\right] \leq e^{-2 p(\epsilon / 3)^{2}}
$$

Plugging in our value for $p$ results in $e^{-2 p(\epsilon / 3)^{2}} \leq \frac{\delta}{2 n}$

## Sample and runtime complexity

Lastly, we need to make sure that this algorithm uses time and sample size expected polynomial in all relevant parameters: $\epsilon, \delta, n, m$.
First, some definitions:
(1) Let $T_{i}$ be the runningtime of $L^{\prime}$ on round $i$. Note that each $T_{i}$ can be different, since $L^{\prime}$ has time complexity which can depend on the number of terms $m$.
(2) Let $P_{i}$ be the probability that we reach round $i$
(3) Let $k$ be the smallest natural number for which $2^{k} \geq m$. I.e. $k=\lceil\log (m)\rceil$

First, we observe that the time it takes to complete the runs for which $m_{i}^{\prime}<m$ (where $m_{i}^{\prime}=2^{i}$ is the $m$ we pass to $L$ at round $i$ ) is polynomial.

$$
\begin{aligned}
\sum_{i=0}^{k-1} P_{i} T_{i} & \leq \sum_{i=0}^{k-1} T_{i} \\
& \leq k \max _{i}\left(T_{i}\right) \\
& \leq\lceil\log (m)\rceil \operatorname{poly}\left(\frac{4}{\epsilon}, 3 * 2^{c}, n, m_{i^{*}}^{\prime}\right) \\
& =\operatorname{poly}\left(\frac{1}{\epsilon}, \frac{1}{\delta}, n, m\right)
\end{aligned}
$$

In the second to last line, $i^{*}=\arg \max _{i}\left(T_{i}\right)$. And we obtain the final expression by absorbing the constants and the $\log (m)$ factor, and recalling that $m_{i}^{*}<m$ by definition.

Second, we want to upperbound the expected amount of time (samples) we spend on rounds with $m_{i}^{\prime}>m$. Note that above, we showed that the runtime is guaranteed polynomial, since in our analysis, we completely ignored the $P_{i}$ 's. However, now the $P_{i}$ 's play an important role. Since there's always a chance that we keep getting bad test sets, we might try all $m^{\prime}$ up to $3^{n}$. We need to ensure that the probability of reaching large $m_{i}^{\prime}$ (relative to the true $m$ ) is "small enough." More formally:
Define $U_{j}=T_{j+k}$ be the running time for the $j$ th round for which $m^{\prime}>m$. Observe:

$$
\begin{aligned}
U_{j} & =T_{j+k} \\
& =T_{k} * \frac{T_{j+k}}{T_{k}} \\
& \leq T_{k} *\left(\frac{m_{j+k}^{\prime}}{m_{k}^{\prime}}\right)^{c * j} \\
& =T_{k} * 2^{c * j}
\end{aligned}
$$

Recalling that $T_{k}$ is polynomial in all problem parameters, it suffices if the probability to reach the $j$ th extra round is bounded by $\frac{1}{2^{c j}}$.
Define $Q_{j}=P_{j+k}$ to be the probability we reach the $j$ th round after $m^{\prime}>m$. And define $R_{i}$ to be the probability that, given that we've reach round $i$, we reject the $i$ th hypothesis and keep going. Note that:

$$
Q_{j}=\prod_{i=0}^{j+k-1} R_{i} \leq \prod_{i=k}^{j+k-1} R_{i}
$$

The idea is that, when $m_{i}^{\prime}>m$, the probability that we accept is high (and so the probability that we reject is low). The only times we reject are if: (a) $L$ fails to produce an $\epsilon / 4$ accurate hypothesis, or (b) $L$ produces an accurate hypothesis, but the sample we draw is unrepresentative. Define $R_{i}^{a}, R_{i}^{b}$ to be the probabilities of these respective events, where from now on, we assume $i>k$.
$R_{i}^{a}$ follows from the $\delta$ parameter we passed to $L$, and the fact that $L$ is a PAC learning algo:

$$
R_{i}^{a} \leq \delta_{i}=\frac{1}{3 * 2^{c}}
$$

$R_{i}^{b}$ follows from the exact same Chernoff bound analysis as for the correctness analysis from before, with the desired probability bound of $\frac{\delta}{2 n}$ replaced with $\frac{1}{3 * 2^{c}}$.

$$
R_{i}^{b} \leq \delta_{i}=\frac{1}{3 * 2^{c}}
$$

Alright... we can finally start piecing things back together. Applying union bound, we get:

$$
R_{i} \leq R_{i}^{a}+R_{i}^{b}=\frac{2}{3 * 2^{c}}<\frac{1}{2^{c}}
$$

Plugging back into the $Q_{j}$ s:

$$
Q_{j} \leq \prod_{i=k}^{j+k-1} R_{i}<\frac{1}{2^{c j}}
$$

And finally, we can use this to compute the expected runningtime above $m$ :

$$
\begin{align*}
\sum_{i=k}^{n \log 3} P_{i} T_{i} & =\sum_{j=0}^{n \log 3-k} Q_{j} U_{j}  \tag{1}\\
& <\sum_{j=0}^{2 n} T_{k} * 2^{c j} * \frac{1}{2^{c j}}  \tag{2}\\
& =2 n T_{k}  \tag{3}\\
& =\operatorname{poly}\left(\frac{1}{\epsilon}, \frac{1}{\delta}, n, m\right) \tag{4}
\end{align*}
$$

Finally... we can conclude that the expected time complexity of this process is polynomial in all relevant parameters. Note that we only showed this for time complexity, but you can derive the same result for sample complexity by following the above analysis.

Notice that we didn't use much about the fact that we were specifically learning DNFs. The only part we somewhat used that fact was in stating that $m$ was bounded by $3^{n}$. Because of the generality of the proof, we can observe that the same overall algorithm can be used to turn any PAC learning algorithm which depends on a bounded integral parameter into an (expected polynomial) PAC learning algorithm which does not require the parameter as input.

